

## Topic 4: Generalized Least Squares

### Generalized Least Squares

(Alexander Aitken, 1935)

- In the present context, (Ordinary) LS ignores some important information, and we'd anticipate that this will result in a loss of efficiency when estimating  $\boldsymbol{\beta}$ .
- Let's see how to obtain the fully efficient (linear unbiased) estimator.
- Recall that  $V(\boldsymbol{\varepsilon}) = E[\boldsymbol{\varepsilon}\boldsymbol{\varepsilon}'] = \Sigma = \sigma^2\Omega$ .
- Generally,  $\Omega$  will be *unknown*. However, to begin with, let's consider the case where it is actually *known*.
- Clearly,  $\Omega$  must be *symmetric*, as it is a covariance matrix.
- Suppose that  $\Omega$  is also *positive-definite*.
- Then,  $\Omega^{-1}$  is also positive-definite, and so there exists a *non-singular* matrix,  $P$ , such that  $\Omega^{-1} = P'P$ .
- Our model is:

$$\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon} \quad ; \quad \boldsymbol{\varepsilon} \sim [0, \sigma^2\Omega]$$

- Pre-multiply the equation by  $P$ :

$$P\mathbf{y} = PX\boldsymbol{\beta} + P\boldsymbol{\varepsilon}$$

or,

$$\mathbf{y}^* = X^*\boldsymbol{\beta} + \boldsymbol{\varepsilon}^* \quad ; \quad \text{say}$$

- Now,  $\Omega$  is non-random, so  $P$  is also non-random.
- So,  $E[\boldsymbol{\varepsilon}^*] = E[P\boldsymbol{\varepsilon}] = P E[\boldsymbol{\varepsilon}] = \mathbf{0}$ .

- And  $V[\boldsymbol{\varepsilon}^*] = V[P\boldsymbol{\varepsilon}]$ 

$$= PV(\boldsymbol{\varepsilon})P'$$

$$= P(\sigma^2\Omega)P' = \sigma^2 P\Omega P'$$

- Note that  $P\Omega P' = P(\Omega^{-1})^{-1}P'$

$$= P(P'P)^{-1}P'$$

$$= PP^{-1}(P')^{-1}P' = I$$

- (Because  $P$  is both *square and non-singular*.)

- So,  $E[\boldsymbol{\varepsilon}^*] = \mathbf{0}$  and  $V[\boldsymbol{\varepsilon}^*] = \sigma^2 I$ .
- The transformed model,  $\mathbf{y}^* = X^* \boldsymbol{\beta} + \boldsymbol{\varepsilon}^*$ , has an error-term that satisfies the *usual assumptions*. **In particular, it has a scalar covariance matrix.**
- So, if we apply (Ordinary) Least Squares to the model,  $\mathbf{y}^* = X^* \boldsymbol{\beta} + \boldsymbol{\varepsilon}^*$ , we'll get the BLU estimator of  $\boldsymbol{\beta}$ , by the Gauss-Markhov Theorem.
- We call this the **Generalized Least Squares Estimator** of  $\boldsymbol{\beta}$ .
- The formula for this estimator is readily determined:

$$\begin{aligned}\widehat{\boldsymbol{\beta}} &= [X^{*'} X^*]^{-1} X^{*'} \mathbf{y}^* \\ &= [(PX)'(PX)]^{-1} (PX)'(P\mathbf{y}) \\ &= [X' P' P X]^{-1} X' P' P \mathbf{y} \\ &= [X' \Omega^{-1} X]^{-1} X' \Omega^{-1} \mathbf{y}\end{aligned}$$

- Clearly, because  $E[\boldsymbol{\varepsilon}^*] = \mathbf{0}$  as long as the regressors are non-random, the GLS estimator,  $\widehat{\boldsymbol{\beta}}$  is **unbiased**.
- Moreover, the covariance matrix of the GLS estimator is:

$$\begin{aligned}V(\widehat{\boldsymbol{\beta}}) &= [X' \Omega^{-1} X]^{-1} X' \Omega^{-1} V(\mathbf{y}) \{ [X' \Omega^{-1} X]^{-1} X' \Omega^{-1} \}' \\ &= [X' \Omega^{-1} X]^{-1} X' \Omega^{-1} \sigma^2 \Omega \Omega^{-1} X [X' \Omega^{-1} X]^{-1} \\ &= \sigma^2 [X' \Omega^{-1} X]^{-1}.\end{aligned}$$

- If the errors are Normally distributed, then the full sampling distribution of the GLS estimator of  $\boldsymbol{\beta}$  is:

$$\widehat{\boldsymbol{\beta}} \sim N[\boldsymbol{\beta}, \sigma^2 [X' \Omega^{-1} X]^{-1}]$$

- The GLS estimator is just the OLS estimator, applied to the transformed model, and the latter model satisfies all of the usual conditions.

- So, the *Gauss-Markhov Theorem* applies to the GLS estimator.
- The GLS estimator is BLU for this more general model (with a non-scalar error covariance matrix).
- Note: OLS must be *inefficient* in the present context.

### “Feasible” GLS Estimator

- In order to be able to implement the GLS estimator, in practice, we’re usually going to have to provide a *suitable estimator* of  $\Omega$  (or  $\Sigma$ ).
- Presumably we’ll want to obtain an estimator that is *at least consistent*, and place this into the formula for the GLS estimator, yielding:

$$\tilde{\beta} = [X'\hat{\Omega}^{-1}X]^{-1}X'\hat{\Omega}^{-1}\mathbf{y}$$

- Problem: The  $\Omega$  matrix is  $(n \times n)$ , and it has  $n(n + 1)/2$  *distinct* elements. However, we have only  $n$  observations on the data. This precludes obtaining a consistent estimator.
- We need to constrain the elements of  $\Omega$  in some way.
- In practice, this won’t be a big problem, because usually there will be lots of “structure” on the form of  $\Omega$ .
- Typically, we’ll have  $\Omega = \Omega(\boldsymbol{\theta})$ , where the vector,  $\boldsymbol{\theta}$  has low dimension.

### Example: Heteroskedasticity

Suppose that  $var.(\varepsilon_i) \propto (\theta_1 + \theta_2 z_i) = \sigma^2(\theta_1 + \theta_2 z_i)$

Then,

$$\Omega = \begin{pmatrix} \theta_1 + \theta_2 z_1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \theta_1 + \theta_2 z_n \end{pmatrix}$$

There are just two parameters that have to be estimated, in order to obtain  $\hat{\Omega}$ .

### Example: Autocorrelation

Suppose that the errors follow a *first-order autoregressive process*:

$$\varepsilon_t = \rho\varepsilon_{t-1} + u_t \quad ; \quad u_t \sim N[0, \sigma_u^2] \quad (\text{i.i.d.})$$

Then (for reasons we'll see later),

$$\Omega = \frac{\sigma_u^2}{1-\rho^2} \begin{bmatrix} 1 & \rho & \dots & \rho^{n-1} \\ \rho & 1 & \rho & \rho^{n-2} \\ \vdots & \rho & \ddots & \vdots \\ \rho^{n-1} & \dots & \dots & 1 \end{bmatrix} = \Omega(\rho).$$

- So, typically, we'll just have to estimate a very small number of parameters in order to get an estimator of  $\Omega$ .
- As long as we use a *consistent estimator* for these parameters – the elements of  $\theta$ , this will give us a consistent estimator of  $\Omega$  and of  $\Omega^{-1}$ , by Slutsky's Theorem.
- This in turn, will ensure that our Feasible GLS estimator of  $\beta$  is also *weakly consistent*:

$$\begin{aligned} \text{plim}(\tilde{\beta}) &= \text{plim} \left\{ [X' \hat{\Omega}^{-1} X]^{-1} X' \hat{\Omega}^{-1} \mathbf{y} \right\} \\ &= \text{plim} \left\{ [X' \Omega^{-1} X]^{-1} X' \Omega^{-1} \mathbf{y} \right\} \\ &= \text{plim}(\hat{\beta}) = \beta . \end{aligned}$$

- Also, if  $\hat{\Omega}$  is consistent for  $\Omega$  then  $\tilde{\beta}$  will be *asymptotically efficient*.
- In general, we can say little about the *finite-sample* properties of our feasible GLS estimator.
- Usually it will be *biased*, and the nature of the bias will depend on the form of  $\Omega$  and our choice of  $\hat{\Omega}$ .
- In order to apply either the GLS estimator, or the feasible GLS estimator, we need to know the form of  $\Omega$ .

Typically, this is achieved by postulating various forms, and testing to see if these are supported by the data.