

Assignment #4 Answer Key

$$\begin{aligned} \text{Q.1. a) } L &= p(y_1, \dots, y_n | \beta, \sigma) \\ &= \prod_{i=1}^n p(y_i | \beta, \sigma) \quad ; \text{ given independence} \\ &= \prod_{i=1}^n \left[\frac{1}{\sigma\sqrt{2\pi}} \cdot e^{-\frac{1}{2\sigma^2}(y_i - x_i'\beta)^2} \right] \\ &= \frac{1}{\sigma^n (\sqrt{2\pi})^n} \cdot e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - x_i'\beta)^2} \\ &= \sigma^{-n} (2\pi)^{-n/2} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n [y_i - x_i'\beta]^2} \end{aligned}$$

$$\begin{aligned} \text{b) } \log L &= -n \log \sigma - \frac{n}{2} \log 2\pi - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - x_i'\beta)^2 \\ &= -n \log \sigma - \frac{n}{2} \log 2\pi - \frac{1}{2\sigma^2} (y - X\beta)'(y - X\beta) \end{aligned}$$

$$\begin{aligned} \frac{\partial \log L}{\partial \beta} &= \frac{-1}{2\sigma^2} \cdot \frac{\partial}{\partial \beta} [y'y + \beta'X'X\beta - 2y'X\beta] \\ &= \frac{-1}{2\sigma^2} (2X'X\beta - 2X'y) = 0 \end{aligned}$$

Hence, $(X'X\beta - X'y) = 0 \Rightarrow \hat{\beta} = (X'X)^{-1}X'y$,

and this is just the OLS estimator, which is unbiased under the usual set of assumptions.

c) With OLS, the derivation of b requires:

- (i) Linear model;
- (ii) X has full rank.

(the properties of b then depend on other assumptions).

With the MLE, the derivation of the estimator requires

(i) and (ii) above, and an appropriate assumption about

the distribution of the errors - here, normality was used. The properties of the MLE are the same as those of b , if the same assumptions hold.

$$\begin{aligned} d) \frac{\partial \log L}{\partial \sigma^2} &= \frac{\partial}{\partial \sigma^2} \left[-\frac{n}{2} \log \sigma^2 - \frac{n}{2} \log 2\pi - \frac{1}{2\sigma^2} (y - X\beta)'(y - X\beta) \right] \\ &= -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} (y - X\beta)'(y - X\beta) = 0 \end{aligned}$$

$$\text{So, } \hat{\sigma}^2 = \frac{1}{n} (y - X\hat{\beta})'(y - X\hat{\beta}) = \frac{1}{n} e'e,$$

where $e = \text{MLE residual vector} = \text{OLS residual vector}$

Compare this with $s^2 = \frac{1}{n-k} e'e$, our usual estimator. We know that s^2 is unbiased, so clearly $\hat{\sigma}^2$ is biased.

$$e) E(\hat{\sigma}^2) = \frac{1}{n} E(e'e) = \frac{1}{n} (n-k) \sigma^2 = \left(1 - \frac{k}{n}\right) \sigma^2 \neq \sigma^2 \text{ (Biased)}$$

$$\text{and Bias}(\hat{\sigma}^2) = E(\hat{\sigma}^2) - \sigma^2 = \left(1 - \frac{k}{n}\right) \sigma^2 - \sigma^2 = \left(-\frac{k}{n}\right) \sigma^2.$$

So, Bias is negative. For fixed k , if $n \rightarrow \infty$, then this bias vanishes.